Intake of Various Food Groups and Risk of Breast Cancer: A Systematic Review and Dose-Response Meta-Analysis of Prospective Studies

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ABSTRACT

Despite increasing evidence for the association of food-based dietary patterns with breast cancer risk, knowledge about the shape of the relationship and the quality of meta-evidence are insufficient. We aimed to summarize the associations between food groups and risks of breast cancer. We performed a systematic literature search of the PubMed and Embase databases up to March 2020. We included cohort, case-cohort, nested casecontrol studies, and follow-up studies of randomized controlled trials that investigated the relationship between breast cancer risk and at least 1 of the following food groups: red meat, processed meat, fish, poultry, egg, vegetables, fruit, dairy product (overall, milk, yogurt, and cheese), grains/cereals, nuts, legumes, soy, and sugar-sweetened beverages. Summary risk ratios (RRs) and 95% CIs were estimated using a random-effects model for linear and nonlinear relationships. Inverse linear associations were observed for vegetables (RR per 100 g/d, 0.97; 95% CI, 0.95–0.99), fruit (RR per 100 g/d, 0.97; 95% CI, 0.95–0.99), cheese (RR per 30 g/d, 0.95; 95% CI, 0.91–1.00), and soy (RR per 30 g/d, 0.96; 95% CI, 0.94–0.99), while positive associations were observed for red (RR per 100 g/d, 1.10; 95% CI, 1.03–1.18) and processed meat (RR per 50 g/d, 1.18; 95% CI, 1.04–1.33). None of the other food groups were significantly associated with breast cancer risk. A nonlinear association was observed only for milk, such that the intake of >450 g/d increased the risk, while no association was observed for lower intake amounts. High intakes of vegetables, fruit, cheese, and soy products and low intakes of red and processed meat were associated with lower risks of breast cancer. However, causality cannot be inferred from these statistical correlations. Adv Nutr 2021;12:809–849.

Keywords: breast cancer, hormone receptor-positive, ER/PR-positive, refined grain, whole grain

Introduction

Breast cancer is the most commonly diagnosed cancer among females and the leading cause of cancer-related death in women. In 2018, 2.1 million new breast cancer cases were estimated, accounting for approximately 11.6% of all cancers in the world [\(1\)](#page-37-0), whilst in the same time-period, an estimated 600 000 deaths occurred worldwide in 2018, accounting for 6.6% of deaths from all cancer types [\(2\)](#page-37-1).

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Knowledge of the etiology of breast cancer is still limited [\(2\)](#page-37-1), but a variety of modifiable and nonmodifiable risk factors have been identified. Indeed, race, ethnicity, family history of cancer, and genetic traits have been identified as important nonmodifiable risk factors in epidemiologic studies. However, modifiable risk factors have also been identified, such as increased alcohol consumption, physical inactivity, exogenous hormone uses, and certain female reproductive factors, such as pregnancy and age at first birth [\(3\)](#page-37-2). Importantly, the potential role of diet on the risk of breast cancer has been examined in a large volume of epidemiologic studies; however, the specific associations between numerous specific food groups and breast cancer risks are relatively unclear.

Indeed, multiple systematic reviews and meta-analyses have evaluated the association of single food groups with breast cancer risks, and most of the prior meta-analyses

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Supplemental Tables 1 and 2 and Supplemental Figures 1–19 are available from the "Supplementary data" link in the online posting of the article and from the same link in the online table of contents at [https://academic.oup.com/advances/.](https://academic.oup.com/advances/)

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Abbreviations used: IGF-I, insulin-like growth factor I; RCTs, randomized controlled trial; RR, risk ratio; SSB, sugar-sweetened beverages.

have only compared breast cancer risks in the highest versus lowest intakes of selected food groups. Moreover, multiple systematic reviews have examined dietary patterns (i.e., multiple food groups in combination) and breast cancer risks, finding moderate evidence to indicate that dietary patterns rich in vegetables, fruits, and whole grains and lower in animal-source foods and refined carbohydrates are correlated with decreased risks of postmenopausal breast cancer. The data pertaining to these dietary patterns and premenopausal breast cancer risks follow the same direction, but the evidence remains insufficient since few studies include premenopausal breast cancer [\(4,](#page-37-3) [5\)](#page-37-4). However, the present study seeks to strengthen the field by taking a novel approach to examining individual foods/food groups. Thus, the objective of our comprehensive meta-analysis was to assess the shape of the diet/breast cancer relationship by performing linear and nonlinear does-response analyses. We estimated the summary associations between intake of 13 food groups [as defined by the Schwingshackl et al. methodology [\(6\)](#page-37-5)] and breast cancer risks.

Methods

The protocol of this meta-analysis has been registered in the International Prospective Register of Systematic Reviews (PROSPERO; [www.crd.york.ac.uk/prospero/index.asp;](http://www.crd.york.ac.uk/prospero/index.asp) identifier CRD42019144956). This systematic review was developed based on the standards of the Meta-Analysis of Observational Studies in Epidemiology guidelines $(7).$ $(7).$

Study selection

To be eligible for inclusion, studies were required to: *1*) be of cohort, case-cohort, or nested case-control design, including follow-up studies of randomized controlled trials (RCTs); *2*) provide data on the association between the risk of breast cancer and at least 1 of the following 13 food groups: grains/cereals, vegetables, fruit, eggs, dairy products (overall or milk, yogurt, and cheese), fish, poultry, red meat, processed meat, nuts, legumes, soy product, sugarsweetened beverages (SSB); 3) include participants aged \geq 18 y; and *4*) assess dietary intake at the beginning of the study. When dietary intake was assessed during adolescence or early adulthood, the study was not included in our metaanalysis. If \geq 2 studies were published on the same exposureoutcome pair, we included only the most recent study with the longest follow-up, and thus the greatest number of events. Moreover, studies that only investigated the highest versus lowest categories were excluded. We also excluded studies conducted on micro- and macronutrients (i.e., soy fiber or phytoestrogen), and focused our evaluation on dietary groups. Studies that only assessed cancer recurrence or survivorship as the outcome were excluded, and studies with case-control and cross-sectional designs and RCTs and non-RCTs were excluded. We imposed no limitation or restriction on the geographical location and health status of participants.

Search strategy

Articles published through March 2020 and indexed in PubMed and Embase were searched for prospective studies, based on the above inclusion criteria, with no language restriction. The search terms used as keywords in the search strategy are listed in **Supplemental Table 1**. In addition, the bibliographies of all relevant prior reviews and primary studies identified by the electronic search strategy were scanned for relevant papers.

Data extraction

Our 2 reviewers independently extracted the following information: name of first author, year of publication, country, cohort name, age at entry, menopause status, sample size, total cases, dietary assessment, outcome, outcome assessment, type and quantity of food group, adjustment factors, duration of study, and risk estimate [risk ratios (RRs), HRs, or ORs with their corresponding 95% CIs]. Results for the fully adjusted model were extracted as the preferential data for our analyses. When a study did not report sufficient information for data extraction, we contacted the corresponding author by e-mail at least 2 times, 1 week apart; accordingly, we attained additional data for 2 papers using this method [\(8,](#page-37-7) [9\)](#page-37-8). For the linear dose-response relationship, no studies were excluded because of incomplete data. But for the nonlinear analysis, 9 studies did not report the number of cases in each category and 1 study did not provide data on the amount of dietary intake in each category. Since we could not obtain required data after contacting the corresponding authors, we excluded these studies from the nonlinear analysis.

Risk of bias assessment and quality of evidence

We used the Newcastle-Ottawa Scale to assess the methodological quality of included studies [\(10\)](#page-37-9). We examined 3 main domains—selection, comparability, and outcome to rate the quality of studies. In the selection domain, 4 items were assessed: representativeness of the exposed cohort, selection of the nonexposed cohort, ascertainment of exposure, and demonstration that the outcomes were not present at the start of the study. In the comparability domain, the control of confounders in the design or analysis of the studies was checked. Finally, in the outcome domain, the outcomes ascertainment, duration of follow-up, and adequacy of follow-up of cohorts were considered. If a study received 3–4 stars in the selection domain, 1–2 stars in the comparability domain, and 2–3 stars in the outcome domain, the quality was rated as good. If a study received 2 stars in the selection domain, 1–2 stars in the comparability domain, and 2–3 stars in the outcome domain, the quality was rated as fair. If a study received 0–1 star in the selection domain, 0 stars in the comparability domain, or 0–1 star in the outcome domain, the quality was rated as poor.

The overall quality of the studies included in this metaanalysis was also evaluated by the use of the NutriGrade scoring system [\(11\)](#page-37-10), which comprises the following items: *1*) risk of bias, study quality, and study limitations (0–2 points); *2*) precision (0–1 point); *3*) heterogeneity (0–1 point); *4*)

directness (0–1 point); *5*) publication bias (0–1 point); *6*) funding bias $(0-1 \text{ point})$; *7*) effect size $(+ 2 \text{ points})$; and *8*) dose response $(+1 \text{ point})$. This scoring system recommends 4 categories to define the meta-evidence as high (≥ 8 points), moderate (6–7.99 points), low (4–5.99 points), or very low (0–3.99 points).

Statistical analysis

We used HRs and 95% CIs as the effect sizes for all analyses. The reported RRs or ORs in the primary studies were considered to be equal to HRs. The dose-response meta-analysis was performed using the method proposed by Greenland and Longnecker [\(12\)](#page-37-11) and Orsini et al. [\(13\)](#page-37-12) and consists of 2 parts: linear analysis and nonlinear analysis. Using a random-effects model, we performed a linear dose-response meta-analysis by pooling the HRs for each increment of 100 grams of meat, poultry, fish, fruit, and vegetable intake; 50 grams of processed meat, egg, fruit juice, and legume intake; 200 grams of dairy (as a whole), milk, and yogurt intake; 30 grams of cheese and soy intake; 20 grams of cereals intake; and 28 grams of nut intake.

To assess the nonlinear dose-response relationship, a 2-stage hierarchical regression model was used, in which the difference between category-specific and reference-specific doses, expressed in quadratic terms, was calculated. Then, the dose-response association, considering within- and betweenstudy variances, was estimated through the use of spline transformations. This method requires the distribution of cases and noncases across >3 categories of food groups, using the median value and the adjusted RRs with their 95% CIs for each category of exposure. For the estimation quantity of food consumption, the median intake of each food group was used. If a study reported both the mean and median of the group, we used the median. Only mean intakes were reported in 11 papers, so for these studies the mean intake was used. In instances where the amount of food intake in each category was reported in the closed interval, consumption was considered as the midpoint of the interval. For the open-ended exposure categories, we considered the length of the open-ended interval to be the same as that of the adjacent interval. We set 2-sided statistical significance a priori at *P* < 0.05.

The Q test and the I^2 statistic (with a value of $I^2 > 50\%$ considered to represent potentially important statistical heterogeneity) was used to explore heterogeneity between studies. To discern the source of heterogeneity, we performed subgroup analyses of potential influencing factors, including menopause status, presence of estrogen receptor, followup duration, geographical location, number of cases, and characteristics of the food items (e.g., high- vs. low-fat content or whole vs. refined grain). However, it was not possible to perform subgroup analyses by all of these factors for all of food groups, because in some cases fewer than 2 studies were in a subgroup or the primary studies did not report the results appropriately; for example, for milk, some primary studies reported data separately according to the fat content (low- vs. high-fat intake) for dairy, yogurt, cheese, and meat, while some studies did not report results according to the fat content.

If at least 10 studies were available, we explored potential small-study effects, such as publication bias, by using Egger's test and funnel plots. Stata version 13 software was used to conduct all statistical analyses.

Results

As detailed in **[Figure 1](#page-3-0)**, 7635 records were obtained following the literature search. Of these, 210 articles were potentially relevant for inclusion in the meta-analysis because they reported \geq 1 of the 13 food groups and breast cancer risk in the title or abstract. Finally, the number of studies included in the meta-analysis for each food group were as follows: total meat: 13; red meat: 20; processed meat: 17; poultry: 13; fish: 17; egg: 11; fruit: 15; vegetable: 14; dairy: 10; milk: 13; yogurt: 6; cheese: 10; total cereals (both whole and refined): 14; soy and soy products: 7; nuts: 6; and legumes: 4. The number of studies on SSB was not adequate. The included studies were performed in Asia, Europe, North America, and Australia (1 study), and characteristics of all studies are presented in **[Table 1](#page-4-0)**.

Total meat

From 14 studies, we investigated the association of total meat consumption with breast cancer, where 1 study was excluded [\(14\)](#page-37-13) due to an identical publication with a longer duration being available. Therefore, 13 studies, with 48 590 breast cancer cases, were included in the linear dose-response meta-analysis [\(15–28\)](#page-37-14). Each additional 100-g/d increase of total meat was associated with a small increase in the risk of breast cancer (RR, 1.07; 95% CI, 1.01–1.13; *I* 2, 75.5%; *P*heterogeneity < 0.001; **Supplemental Figure 1**). A subgroup analysis by duration, number of cases, and location indicated that this association persisted only in studies with a duration of <10 y, case numbers of \geq 1000, and studies conducted in Europe; while in studies with a duration of \geq 10 y, case numbers of <1000, and studies conducted in the United States, no association was observed (**[Table 2](#page-26-0)**). Moreover, the difference between premenopausal and postmenopausal status was nonsignificant [\(Table 2\)](#page-26-0).

We found no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.21; $n = 11$ studies; **[Figure 2](#page-30-0)**A).

Red meat

The association of red meat with breast cancer was investigated by 26 articles. We excluded 6 papers because other papers on the same cohort with longer durations were published [\(14,](#page-37-13) [29–33\)](#page-38-0); thus, only the most recent studies with the longest follow-ups were included. These 20 studies [\(15,](#page-37-14) [16,](#page-37-15) [18–22,](#page-37-16) [24,](#page-37-17) [34–43\)](#page-38-1), with 78 267 breast cancer cases, were included in the linear dose-response meta-analysis. Each 100-g/d increase of red meat was associated with a small increase in the risk of breast cancer (RR, 1.10; 95% CI, 1.03–1.18); however, statistically significant heterogeneity was observed in this model (I^2 , 60.2%; *P*-heterogeneity < 0.001; **Supplemental Figure 2**). The observed positive

FIGURE 1 Flowchart of study selection.

associations persisted in additional analyses, stratified by follow-up duration, geographic location, number of cases, and menopausal status [\(Table 2\)](#page-26-0). The subgroup differences were not statistically significant, with the exception of followup duration, which showed a stronger inverse association for studies with a duration of <10 y.

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.24; *n* = 14 studies). The risk of breast cancer increased by approximately 10% with increasing intake of red meat, up to 150 g/d [\(Figure 2B](#page-30-0)).

Processed meat

From 20 papers that investigated the relationship between processed meat and breast cancer, 3 articles [\(14,](#page-37-13) [32,](#page-38-2) [33\)](#page-38-3) were excluded because the same exposure-outcome pair with a longer duration was published. There were 17 studies, with 34 414 breast cancer cases, included in the linear doseresponse meta-analysis [\(15,](#page-37-14) [16,](#page-37-15) [18–20,](#page-37-16) [22,](#page-37-18) [23,](#page-37-19) [25,](#page-37-20) [27,](#page-38-5) [34,](#page-38-1) [37,](#page-38-6) [39–41,](#page-38-7) [44,](#page-38-4) [45\)](#page-38-8). A positive association was observed for each additional 50-g/d increase of processed meat (RR, 1.18; 95% CI, 1.04–1.33; *I* 2, 63.5%; *P*-heterogeneity < 0.001; **Supplemental Figure 3**). Subgroup analyses by menopause

status and follow-up duration indicated no significant association [\(Table 2\)](#page-26-0). However, subgroup analyses by number of cases and geographic location revealed a stronger positive association for studies with case numbers of <1000 and studies conducted in Europe [\(Table 2\)](#page-26-0).

No evidence of a nonlinear dose-response association was detected (*P*-nonlinearity = 0.10; $n = 15$ studies). The risk of breast cancer increased by approximately 10% with an increasing intake of processed meat, up to 50 g/d [\(Figure 2C](#page-30-0)).

Poultry

From 14 studies that investigated the association of poultry with breast cancer, 1 study was excluded [\(32\)](#page-38-2) because a paper on the same cohort with a longer duration was published. Therefore, 13 studies, with 27 445 breast cancer cases, were included in a linear dose-response meta-analysis [\(15,](#page-37-14) [16,](#page-37-15) [19,](#page-37-21) [22–26,](#page-37-18) [36,](#page-38-9) [37,](#page-38-6) [39,](#page-38-7) [40,](#page-38-10) [46\)](#page-38-11). No association was observed for each 100-g/d increase of poultry (RR, 0.97; 95% CI, 0.91–1.03; *I* 2, 22.9%; *P*-heterogeneity, 0.21; **Supplemental Figure 4**). Subgroup analyses by menopause status, follow-up duration, and geographic location indicated no significant as-sociation in the subgroups [\(Table 2\)](#page-26-0). A subgroup analysis by TABLE 1 General study characteristics of the included studies investigating the association between various food groups and risk of breast cancer **TABLE 1** General study characteristics of the included studies investigating the association between various food groups and risk of breast cancer

multivitamin use

(Continued)

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sweets, and eggs

(Continued) (Continued)

status

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and baseline MHT

(Continued)

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Age, study area, family

Age, study area, family

Adjusted covariate

utcome

first birth, parity, use of walking time, BMI, and first birth, parity, use of consumption of green duration of HRT, intake walking time, BMI, and duration of HRT, intake consumption of green breast tumor surgery, education, use of HRT, education, use of HRT, increase since age 18, increase since age 18, activity; MHT; alcohol combined age at first Age, duration, baseline breast tumor surgery, consumption; intakes consumption; intakes of red meat and total combined age at first education, maternal activity; MHT; alcohol of red meat and total Age, El, alcohol intake, menopause, age at Age, duration, baseline practice, and use of of alcohol, and BMI education, maternal practice, and use of Cancer registry Age, EI, alcohol intake, menopause, age at of alcohol, and BMI education; marital status; BMI; family education; marital status; BMI; family history of cancer; vigorous physical energy; smoking; previous benign history of breast history of breast history of cancer; vigorous physical energy; smoking; leafy vegetables, leafy vegetables, values of parity, height, weight, level of physical activity, years of cancer,
mammography history of breast previous benign history of breast activity, years of mammography HRT, smoking, values of parity, level of physical cancer, age at height, weight, cancer, age at HRT, smoking, Race/ethnicity, Race/ethnicity, total EI OCP members to state members to state cancer registries cancer registries National Death and to the US National Death and to the US Cancer registry Linking cohort cancer Linking cohort Danish cancer try Index Validated FFQ (260), UK and (260), UK and Validated FFQ Validated FFQ NIH-AARP, USA 198 903 50–71 90 5856 Dairy Validated FFQ Sweden (FFQ and a record) 7-day (124) Milk, egg, red meat, milk, processed meat, milk, 319 826 51 NM 7119 Milk, egg, red poultry, cheese meat, Dairy 7119 5856

Pala et al., 2009 [\(40\)](#page-38-10) Prospective

Pala et al., 2009 (40)

EPIC, European countries

countries

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cohort (8.6)

cohort (8.6) Prospective

Park et al., 2009 [\(73\)](#page-39-9) Prospective

Park et al., 2009 (73)

Prospective
cohort (7)

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 $50 - 71$

198903

NIH-AARP, USA

birth and number of children,age at menopause; and intake of fat

menopause; and

intake of fat

children,age at

birth and number of

(Continued) (Continued)

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TABLE 2 (Continued)

TABLE 2 (Continued)

+, positive/present; −, negative/not present; Er, estrogen receptor; Pr, progesterone receptor; RR, risk ratio.

number of cases revealed a stronger positive association for studies with case numbers of <1000. There was no evidence of a nonlinear dose-response association (*P*-nonlinearity = 0.08; $n = 10$ studies; [Figure 2D](#page-30-0)).

Fish

From 18 studies that investigated the association of fish with breast cancer, 1 study was excluded [\(32\)](#page-38-2) because a study on the same cohort with a longer duration was published. Thus, 17 studies, with 28 818 breast cancer cases, were included in a linear dose-response meta-analysis [\(15,](#page-37-14) [16,](#page-37-15) [18,](#page-37-16) [19,](#page-37-21) [23,](#page-37-19) [24,](#page-37-17) [26,](#page-37-25) [35,](#page-38-15) [36,](#page-38-9) [39,](#page-38-7) [42,](#page-38-35) [44,](#page-38-4) [46,](#page-38-11) [53,](#page-38-19) [57,](#page-38-23) [65,](#page-39-1) [80\)](#page-39-16). No association was observed for each additional 100-g/d increase of fish (RR, 1.0; 95% CI, 0.93–1.08; *I* 2, 22.6%; *P*-heterogeneity, 0.19; **Supplemental Figure 5**). Subgroup analyses by menopause status, follow-up duration, number of cases, and geographic location indicated no significant association in the subgroups [\(Table 2\)](#page-26-0).

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.39; $n = 11$ studies). The risk of breast cancer increased by approximately 10% with increasing intake of fish, up to 110 g/d [\(Figure 2E](#page-30-0)).

FIGURE 2 Nonlinear dose-response relationship between daily intakes of (A) total meat, (B) red meat, (C) processed meat, (D) poultry, (E) fish, and (F) egg and risk of breast cancer. $RR =$ risk ratio.

Egg

There were 11 studies, with 53 310 breast cancer cases, included in the linear dose-response meta-analysis of the association between egg intake and breast cancer risk [\(15,](#page-37-14) [17,](#page-37-22) [19,](#page-37-21) [21,](#page-37-24) [23,](#page-37-19) [24,](#page-37-17) [36,](#page-38-9) [39,](#page-38-7) [40,](#page-38-10) [42,](#page-38-35) [44\)](#page-38-4). No association was found for each additional 50-g/d increase of egg (RR, 1.03; 95% CI, 0.96–1.12;*I* 2, 48.8%; *P*-heterogeneity, 0.03; **Supplemental Figure 6**). Subgroup analyses by menopause status, follow-up duration, geographic location, and number of cases indicated no significant difference between the subgroups [\(Table 2\)](#page-26-0).

Although a nonlinear dose-response association was detected, the shape of the curve did not provide any valuable information (*P*-nonlinearity, 0.03; $n = 7$ studies; [Figure 2F](#page-30-0)).

Fruit

Only 1 study assessed the association between combined fruit and vegetable intake and breast cancer, so it was not included in the analyses [\(87\)](#page-39-23).

From 21 papers that investigated the relationship between fruit and breast cancer, 6 were excluded because papers on the same cohorts with longer durations were published [\(31,](#page-38-33) [72,](#page-39-8) [88–91\)](#page-39-24). However, 4 of these studies were included in the subgroup analyses according to menopause status [\(31,](#page-38-33) [72\)](#page-39-8) and presence of estrogen receptor [\(89,](#page-39-25) [90\)](#page-39-26), since the updated papers on the same studies did not report the results according to these factors [\(52,](#page-38-18) [56\)](#page-38-22).

There were 15 studies, with 7071 breast cancer cases, included in the linear dose-response meta-analysis [\(](#page-38-13)[15,](#page-37-14) [21,](#page-37-24) [24,](#page-37-17) [29,](#page-38-0) [35,](#page-38-15) [44,](#page-38-4) [48,](#page-38-13) [49,](#page-38-14) [52,](#page-38-18) [56,](#page-38-22) [59,](#page-38-25) [74,](#page-39-10) [75,](#page-39-11) [78,](#page-39-14) [81,](#page-39-17) [83\)](#page-39-19). A small, inverse association was observed for each additional 100 g/d increase of fruit (RR, 0.97; 95% CI, 0.95–0.99; *I* 2, 66.5%; *P*-heterogeneity < 0.001; **Supplemental Figure 7**). Also, in subgroup analyses by geographic location, number of cases, menopause status, and presence of estrogen progesterone receptor, no significant difference was found between the subgroups [\(Table 2\)](#page-26-0).

FIGURE 3 Nonlinear dose-response relationship between daily intakes of (A) fruit, (B) juice, (C) vegetables, and (D) fruits and vegetables and risk of breast cancer. $RR =$ risk ratio.

There was no evidence of a nonlinear dose-response association detected (*P*-nonlinearity, 0.20; $n = 7$ studies; **[Figure 3](#page-31-0)**A).

Fruit juice

There were 6 studies, with 4463 breast cancer cases, included in the linear dose-response meta-analysis of fruit juice intake and breast cancer risk [\(15,](#page-37-14) [56,](#page-38-22) [58,](#page-38-24) [62,](#page-38-29) [67,](#page-39-3) [78\)](#page-39-14). No significant association was observed for each additional 50-g/d increase of fruit juice (RR, 1.0; 95% CI, 0.99–1.01; *I* 2, 26.7%; *P*heterogeneity, 0.23; **Supplemental Figure 8**).

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.78; $n = 5$ studies; [Figure 3B](#page-31-0)).

Vegetable

From 18 papers that investigated the relationship between vegetable intake and breast cancer, 4 were excluded because papers on the same cohorts with longer durations were published [\(31,](#page-38-33) [72,](#page-39-8) [88,](#page-39-24) [91\)](#page-39-27); however, 1 of these 4 studies [\(72\)](#page-39-8) was included in a subgroup analysis since the updated paper on the same cohort [\(52\)](#page-38-18) did not report the results according to menopause status.

There were 14 studies, with 54 845 breast cancer cases, included in the linear dose-response meta-analysis [\(](#page-38-13)[15,](#page-37-14) [21,](#page-37-24) [24,](#page-37-17) [29,](#page-38-0) [48,](#page-38-13) [49,](#page-38-14) [52,](#page-38-18) [56,](#page-38-22) [59,](#page-38-25) [74,](#page-39-10) [75,](#page-39-11) [78,](#page-39-14) [81,](#page-39-17) [83\)](#page-39-19). A small, inverse association was observed for each additional 100-g/d increase of vegetable intake (RR, 0.97; 95% CI, 0.953–0.995; *I* 2, 55.8%; *P*-heterogeneity, 0.006; **Supplemental Figure**

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9). Subgroup analyses by menopause status, presence of estrogen receptor, geographical location, and number of cases indicated no significant difference in the effect sizes between the subgroups [\(Table 2\)](#page-26-0). The follow-up durations of all studies, except 1, were longer than 10 y.

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.37; $n = 6$, studies; [Figure 3C](#page-31-0)).

Fruit and vegetable

There were 8 studies with breast cancer cases included in the linear dose-response meta-analysis of fruit and vegetable intake [\(24,](#page-37-17) [48,](#page-38-13) [52,](#page-38-18) [56,](#page-38-22) [58,](#page-38-24) [78,](#page-39-14) [81,](#page-39-17) [87\)](#page-39-23). A small, inverse association was observed for each additional 100-g/d increase of fruit and vegetable intake (RR, 0.98; 95% CI, 0.97–0.996; *I* 2, 54.7%; *P*-heterogeneity, 0.03; **Supplemental Figure 10**). Subgroup analyses by geographical location and number of cases indicated no significant difference between the subgroups [\(Table 2\)](#page-26-0). Follow up durations of all studies, except 1, were longer than 10 y.

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.67; $n = 8$ studies; [Figure 3D](#page-31-0)).

Dairy

There were 10 studies, with 16 175 breast cancer cases [\(15,](#page-37-14) [24,](#page-37-17) [26,](#page-37-25) [31,](#page-38-33) [35,](#page-38-15) [55,](#page-38-21) [63,](#page-38-30) [66,](#page-39-2) [73,](#page-39-9) [76\)](#page-39-12), that reported the association of dairy intake as a whole with breast cancer risk. Studies that assessed dairy products separately were not included in this category. A linear dose-response meta-analysis indicated no

FIGURE 4 Nonlinear dose-response relationship between daily intakes of (A) dairy, (B) milk, (C) yogurt, (D) cheese, (E) refined grains, and (F) soy product and risk of breast cancer. $RR = risk$ ratio.

significant association for each additional 200-g/d increase of dairy intake (RR, 0.97; 95% CI, 0.95–1.003; *I* 2, 55.6%; *P*-heterogeneity, 0.02; **Supplemental Figure 11**). Subgroup analyses by menopause status and number of cases indicated no significant difference between the subgroups [\(Table 2\)](#page-26-0). Geographic location and follow-up duration were the sources of heterogenity [\(Table 2\)](#page-26-0).

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.83; $n = 8$ studies; **[Figure 4](#page-32-0)**A).

Milk

There were 13 studies, with 47 729 breast cancer cases, included in the linear dose-response meta-analysis of milk intake and breast cancer risk [\(8,](#page-37-7) [17,](#page-37-22) [18,](#page-37-16) [21,](#page-37-24) [23,](#page-37-19) [39,](#page-38-7) [40,](#page-38-10) [42,](#page-38-35) [45,](#page-38-8) [63,](#page-38-30) [64,](#page-39-0) [68,](#page-39-4) [77\)](#page-39-13). No significant association was observed for each additional 200-g/d increase of milk intake (RR,

0.99; 95% CI, 0.96–1.02; *I* 2, 52.3%; *P*-heterogeneity, 0.01; **Supplemental Figure 12**). Subgroup analyses by menopause status, follow-up duration, geographical location, type of milk, and number of cases indicated no significant difference between the subgroups [\(Table 2\)](#page-26-0).

We found evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.04; $n = 12$ studies). The association of milk intake with breast cancer risk was not significant for intakes of up to 450 g/d, but in amounts greater than 450 g/d, up to 1300 g/d, the risk increased by approximately 30% [\(Figure 4B](#page-32-0)).

Yogurt

There were 6 studies, with 28 291 breast cancer cases, included in the linear dose-response meta-analysis of yogurt intake and breast cancer risk [\(8,](#page-37-7) [21,](#page-37-24) [39,](#page-38-7) [42,](#page-38-35) [64,](#page-39-0) [92\)](#page-39-28). No significant association was observed for each additional 200-g/d increase of yogurt intake (RR, 0.91; 95% CI, 0.79–1.05; *I* 2, 71.5%; *P*-heterogeneity, 0.004; **Supplemental Figure 13**).

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.06; $n = 4$ studies). The risk of breast cancer decreased by approximately 7.5% with an increasing intake of yogurt, up to 100 g/d [\(Figure 4C](#page-32-0)).

Cheese

There were 10 studies, with 39 703 breast cancer cases, included in the linear dose-response meta-analysis of cheese intake and breast cancer risk [\(8,](#page-37-7) [18,](#page-37-16) [21,](#page-37-24) [23,](#page-37-19) [39,](#page-38-7) [40,](#page-38-10) [42,](#page-38-35) [45,](#page-38-8) [63,](#page-38-30) [64\)](#page-39-0). A small, inverse association was observed for each additional 30-g/d increase of cheese intake (RR, 0.95; 95% CI, 0.91–0.996; *I* 2, 75.1%; *P*-heterogeneity < 0.001; **Supplemental Figure 14**). Subgroup analyses for menopause status, follow-up duration, and geographical location were not statistically significant [\(Table 2\)](#page-26-0). A subgroup analysis for the number of cases revealed a stronger inverse association for studies with case numbers of \langle 1000 [\(Table 2\)](#page-26-0).

There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.07; $n = 9$ studies; [Figure 4D](#page-32-0)).

Total cereals

There were 14 studies, with 16 857 breast cancer cases, included in the linear dose-response meta-analysis of total cereal intake and breast cancer risk [\(9,](#page-37-8) [15,](#page-37-14) [24,](#page-37-17) [31,](#page-38-33) [35,](#page-38-15) [44,](#page-38-4) [51,](#page-38-17) [60,](#page-38-26) [69,](#page-39-5) [70,](#page-39-6) [74,](#page-39-10) [78,](#page-39-14) [86,](#page-39-22) [90\)](#page-39-26). No significant association was observed for each additional 20-g/d increase of total cereal intake (RR, 1.0; 95% CI, 0.99–1.01;*I* 2, 17.4%; *P*-heterogeneity = 0.26; **Supplemental Figure 15**). In subgroup analyses by menopause status, follow-up duration, geographic location, number of cases, and refined- vs. whole-grain cereal intake, no significant difference was found in the effect sizes between the subgroups [\(Table 2\)](#page-26-0). Although evidence of a nonlinear dose-response association was detected (*P*-nonlinearity 0.04; $n = 7$ studies), the shape of the curve did not yield any valuable information [\(Figure 4E](#page-32-0)).

Soy and soy products

There were 7 studies, with 4055 breast cancer cases, included in the linear dose-response meta-analysis of soy and soy product intake and breast cancer risk [\(15,](#page-37-14) [24,](#page-37-17) [44,](#page-38-4) [49,](#page-38-14) [69,](#page-39-5) [71,](#page-39-7) [84\)](#page-39-20). A significant association was observed for each additional 30-g/d increase of soy and/or soy product intake (RR, 0.965; 95% CI, 0.94–0.99; *I* 2, 0.0%; *P*-heterogeneity = 0.64; **Supplemental Figure 16**). There was no evidence of a nonlinear dose-response association (*P*-nonlinearity, 0.87; *n* $=$ 5 studies; [Figure 4F](#page-32-0)).

Nuts

There were 6 studies, with 9219 breast cancer cases, included in the linear dose-response meta-analysis of nut intake and breast cancer risk [\(15,](#page-37-14) [36,](#page-38-9) [58,](#page-38-24) [67,](#page-39-3) [78,](#page-39-14) [82\)](#page-39-18). No significant association was observed for each additional 28-g/d increase

Legumes

There were 4 studies that investigated the association of legumes, besides soy, with breast cancer risk [\(15,](#page-37-14) [24,](#page-37-17) [47,](#page-38-12) [58\)](#page-38-24). No significant association was observed for each additional 50-g/d increase of legume intake (RR, 0.95; 95% CI, 0.87– 1.05; *I* 2, 32.1%; *P*-heterogeneity, 0.22).

Publication bias

Based on Egger's test, publication bias was evident only for total meat $(P = 0.007)$, red meat $(P = 0.002)$, and fish (*P* = 0.03) intakes, and their funnel plots (**Supplemental Figure 18**A, B, and E) were asymmetric. There was no publication bias and the associated funnel plots were symmetrical for processed meat (Supplemental Figure 18C), poultry (Supplemental Figure 18D), fruit (**Supplemental Figure 19**A), vegetable (Supplemental Figure 19B), dairy (Supplemental Figure 19C), milk (Supplemental Figure 19D), cheese (Supplemental Figure 19E), and cereal (Supplemental Figure 19F) intakes.

Data quality

The quality of most of the studies was classified as good, while 13 studies were classified as being of fair quality [\(23–26,](#page-37-19) [34,](#page-38-1) [37,](#page-38-6) [62,](#page-38-29) [64,](#page-39-0) [70,](#page-39-6) [71,](#page-39-7) [75,](#page-39-11) [77,](#page-39-13) [83\)](#page-39-19) and 2 studies were classified as being of poor quality [\(44,](#page-38-4) [74\)](#page-39-10) (**Supplemental Table 2**). To discern whether study quality had an effect on the results, we excluded the studies rated as being of fair or poor quality from the analysis; no statistically significant changes were seen, except for in the analysis of total meat. After excluding studies with fair quality, the association of total meat intake with the risk of breast cancer was not significant.

Additionally, the NutriGrade meta-evidence rating indicated moderate confidence in the effect estimates for all of the food categories, except poultry, fish, cereals, and legumes, which had low confidence ratings (**[Table 3](#page-34-0)**).

Discussion

In the present systematic review and meta-analyses, the associations of preselected foods and food groups—total meat, red meat, poultry, fish, processed meat, egg, fruits, vegetables, dairy, milk, yogurt, cheese, grains, soybeans, nuts, and legumes—and the risk of breast cancer were evaluated using data reported within and across prospective studies. We identified decreased risks of breast cancer with increased intakes of fruits, vegetables, soybeans, and cheese, and there was a positive association between red meat and processed meat consumption and the risk of breast cancer. No linear dose-response associations were observed for egg, dairy, milk, yogurt, grain, nut, and legume intakes and breast cancer risks, whilst a nonlinear dose-response association was observed for milk intake. We observed moderate confidence in the effect estimates for all food items, except poultry, fish, cereal, and legumes, which had low confidence ratings.

TABLE 3 NutriGrade assessment of confidence in estimate effect of studies evaluated the association between various food groups and risk of breast cancer **TABLE 3** NutriGrade assessment of confidence in estimate effect of studies evaluated the association between various food groups and risk of breast cancer

1Risk of bias was based on the Newcastle-Ottawa Scale, where ≥7 = 2 points; 4–6.9 = 1 point; and 0–3.9 = 0 points.

nas or us a was based on the reweasties of the second water in the null value; precision is 0 points if the number of events <500 or number of events ≥500, but 95% Cl includes file, Cl includes RR of 1.0) and 95% Cl fails
 Precision is 1 point if the number of events =500 and the 95% Cl excludes the null value; precision is 0 points if the number of events <500 or number of events ≥500, but 95% Cl includes RR of 1.0) and 95% Cl falis to exclude an important benefit (RR of 0.8) or harm (RR of 1.2). to exclude an important benefit (RR of 0.8) or harm (RR of 1.2).

Mhen A was <40% or A was ≥40% but the source of heterogeneity was found by subgroup analysis 1 point was assigned; otherwise, 0 points were assigned. 3When I2 was <40% or I2 was ≥40% but the source of heterogeneity was found by subgroup analysis 1 point was assigned; otherwise, 0 points were assigned.

flased on the funnel plots, Egger or Begg's test. For the outcomes with small number of studies (n < 10), the risk of publication bias was not formally assessed. <code>f</code> based on the funnel plots, Egger or Begg's test. For the outcomes with small number of studies (n $<$ 10), the risk of publication bias was not formally assessed.

Fif the RR or HR <0.80-0.50 and >1.20-2.00, respectively, 1 point is assigned and the corresponding test is statistically significant; if the RR or HR <0.50 and >2.00, respectively, 0 point are assigned and the correspond 16 frithe RR or HR <0.80-0.500, respectively, 1 point is assigned and the corresponding test is statistically significant; if the RR or HR <0.50 and >2.00, respectively, 0 points are assigned and the corresponding tes

significant.
"Moderate quality indicates that we are moderately confident in the effect estimate; the effect is likely to be close to the effect, but there is a possibility that it is substantially different. Low quality i Moderate quality indicates that we are moderately confident in the effect simate: the true refect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different. Low qual in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Lifestyle and environmental factors, including diet, are considered as important factors in the prevention of breast cancer [\(93\)](#page-39-29). The International Agency for Research on Cancer reported that red meat and processed meat may be potential carcinogens for humans [\(94\)](#page-39-30); indeed, in the present meta-analysis, the risk of breast cancer increased by 10% for red meat, 7% for total meat, and 18% for processed meat. Similarly, a previous meta-analysis reported a significant, positive association between processed meat consumption and the risk of breast cancer [\(95\)](#page-39-31), but the authors only compared the highest category with the lowest category of red and processed meat consumption.

The carcinogenicity of red meat and processed meat may be attributed to mutagenic compounds, such as polyaromatic hydrocarbons and heterocyclic amines, which are byproducts of cooking red meat at high temperatures [\(96,](#page-39-32) [97\)](#page-39-33). Also, heme iron, fat, and animal sugar molecule Nglycolylneuraminic acid, found in red meat, are posited to potentially increase inflammation, oxidative stress, and tumor formation [\(96\)](#page-39-32), and in some countries, hormone residue of the exogenous hormones used to stimulate the growth of beef cattle has also been suggested as an independent risk factor of breast cancer [\(96\)](#page-39-32).

To ameliorate the cancer risk, fish and poultry represent good substitutes for red meat in the dietary composition. As in the present meta-analysis, poultry and fish had no significant association with the breast cancer risk. Indeed, red meat and poultry differ in their relative percentages of heme iron and saturated fat content. Also, consumption of poultry has been associated with less mutagenic activity, oxidative stress, and DNA damage [\(93\)](#page-39-29).

Breast cancer has a heterogeneous etiology, so, in the present study, subgroup analyses were conducted based on several factors. In a subgroup analysis, the association of red meat, total meat, and processed meat consumption and breast cancer risk was stronger in the studies from Europe. Indeed, this stronger association might be attributed to the fact that breast cancer is the most common cancer type in Europe [\(98\)](#page-39-34), and such differences might be manifest from the prevalence and distribution of known risk factors of breast cancer in European countries [\(98\)](#page-39-34).

Considering the association between red meat and total meat intake and the breast cancer risk, larger and significant effects were seen in follow-up durations of more than 10 y, which might be attributed to the higher number of cancer cases that occurred in longer follow-ups. Also, cumulative effects of risk factors concomitant to increasing age and an increasing number of post-menopause cancer cases in long follow-ups could be considered, notwithstanding the fact that the effect of red meat consumption on the breast cancer risk was not significant in a subgroup analysis of menopausal status. While some studies revealed that the risks of breast cancer following red meat consumption are different in pre- and postmenopausal women [\(29\)](#page-38-0), in the present study such differences were seen for processed meat intake, which had a stronger association with breast cancer risk in postmenopausal women. Regarding case numbers,

A contentious issue in the relationship between diet and breast cancer risk is dairy consumption. In the present study, a null association was seen between dairy product, milk, and yogurt consumption and breast cancer risks in linear dose-response analyses. A positive nonlinear dosedependent association was seen for milk intake, although no association was observed for consumption of less than 450 g/d of milk, but in amounts greater than 450 g/d, the risk of breast cancer increased by approximately 30%, with increasing milk intake up to 1300 g/d. In a subgroup analysis, the association of dairy consumption with breast cancer risk was larger and was significant in follow-ups of <10 y. The results of a previous meta-analysis investigating the association between milk consumption and breast cancer risk did not provide consistent evidence for such an association [\(99\)](#page-40-0). Indeed, in the aforementioned study, the authors only assessed the relationship between the highest versus lowest intakes of milk; moreover, their literature search was limited to PubMed and Chinese biomedicine databases up to 2009 [\(99\)](#page-40-0), so a number of studies were missing. The association between milk consumption and breast cancer risk might be related to the presence of fat-soluble hormones in the milk, which come from pregnant cows, leading to an increased risk of hormone-dependent cancers, such as breast, ovarian, and corpus uteri cancers [\(100\)](#page-40-1). Milk consumption is among the most important routes of human exposure to estrogens; in fact, milk is considered as the predominant source of animalderived estrogens in the human diet, accounting for 60–80% of the estrogens consumed [\(100\)](#page-40-1). Moreover, milk contains insulin-like growth factor I (IGF-I), which stimulates cell proliferation and neoplasm formation [\(101\)](#page-40-2). The association between milk consumption and IGF-I tumorigenesis was only suggested for milk and not for other dairy products [\(101\)](#page-40-2). Indeed, considering these mechanisms, it is noteworthy to mention that previous meta-analyses showed significant, positive associations between milk consumption and reproductive cancers, such as ovarian cancer [\(102\)](#page-40-3) and prostate cancer [\(101\)](#page-40-2).

In the present study, a small, inverse association was seen between cheese consumption and breast cancer risk, such that the breast cancer risk decreased by 5%. In a subgroup analysis, this association was stronger in studies from Europe. To the best of our knowledge, no meta-analysis has been conducted regarding the association of cheese consumption with breast cancer risk. Cheese is a good dietary source of proteins; several vitamins, such as A, B6, B12, D, and K; and minerals, including calcium, iodine, magnesium, potassium, phosphorus, and zinc. The ameliorative effects of cheese on the breast cancer risk might be because cheese consumption is representative of a relatively healthy diet [\(103\)](#page-40-4). Also, desaturase inhibitors in cheese, which inhibit triglyceride synthesis, reduce the pathogenic effects of fat [\(102\)](#page-40-3), whilst cancer-protective properties of fermented products, such as cheese, might be attributed to live microorganisms acting as probiotics [\(102,](#page-40-3) [104\)](#page-40-5).

In our study, an inverse association was detected between fruit and vegetable consumption and breast cancer risk, such that the risk of breast cancer decreased by 3% for increased fruit intake, by 4% for increased vegetable intake, and by 2% for increased combined fruit and vegetable consumption. The most recent meta-analysis on the associations of fruit and vegetable intakes with breast cancer risk, which was published in 2012, reported that high intakes of fruits alone and of fruits and vegetables combined, but not of vegetables alone, were associated with weak reductions in the breast cancer risk [\(105\)](#page-40-6). There were 9 studies included in this previous study (in a linear dose-response analysis), while our results are based on 15 studies. Anti-cancer properties of fruits and vegetables are possibly due to their high content of antioxidant nutrients, including fiber, vitamins C and E, carotenoids, and other bioactive substances [\(56\)](#page-38-22). In the present study, a subgroup analysis revealed a larger and significant association between vegetable intake and estrogen receptor–negative breast cancer. This might be due to the dominant role of hormonal exposure and hormone-related factors in the etiology of estrogen receptor–positive tumors [\(56,](#page-38-22) [106\)](#page-40-7). Moreover, vegetables contain phytochemical compounds that can reduce the levels of epidermal growth factor receptor, nuclear factor kappa B, and cyclin E, which may, in turn, reduce the risk of developing estrogen receptor– negative breast cancer [\(106\)](#page-40-7).

Besides the antioxidant content of fruits and vegetables, high fiber intake has been shown to interfere with bile acids and decrease estrogen deconjugation, leading to increased fecal excretion of estrogen and reduced plasma concentration of this hormone [\(107\)](#page-40-8). This explanation might also pertain to the significant differences that were seen in a subgroup analysis between refined- and whole-grain consumption in the present study, while a null association was observed between overall grain consumption and breast cancer risk. Also, high glycemic-index foods were shown to be associated with higher insulin levels, and the insulin–IGF-I axis has been shown to be directly associated with cancer promotion [\(108\)](#page-40-9). Similar findings were reported by a previous metaanalysis on whole-grain intake and breast cancer risk, suggesting that intermediate and high intake levels of whole grains were associated with modest reductions of breast cancer risks, but this inverse association was only observed in case-control and not cohort studies [\(109\)](#page-40-10), and the mentioned meta-analysis only assessed studies published specifically on whole grains.

In the present study, soybean consumption was associated with a 3.5% reduction in breast cancer risk. In previous metaanalyses of the association of soy intake and breast cancer, most of the included studies assessed the soy isoflavones. We did not include these studies, since the aim of our study was to investigate the association of food groups, rather than the food component, with breast cancer. There were 2 recent meta-analyses on soy isoflavones that indicated a reduction of breast cancer risk with soy intake, especially in larger amounts. Accordingly, 1 of these studies indicated that high versus low consumption of soy was associated with a lower risk of breast cancer ($n = 6$ studies), while moderate versus low intake of soy did not significantly affect the breast cancer risk ($n = 4$ studies) [\(110\)](#page-40-11). Another study, which evaluated the risk in Chinese women, revealed that every 10 mg/day of soy isoflavone intake was associated with a 3% reduction in breast cancer risk [\(111\)](#page-40-12). Indeed, the results of previous studies indicate that menopause status may influence the association of soy consumption and breast cancer [\(112\)](#page-40-13); however, because of the small number of studies included in our meta-analysis, we could not conduct a subgroup analysis based on menopause status.

The novelties of our study compared to previously published meta-analyses on single food groups [\(95,](#page-39-31) [99,](#page-40-0) [105,](#page-40-6) [109,](#page-40-10) [113\)](#page-40-14) have been explicated above. The only meta-analysis considering several food groups was conducted by Wu et al. in 2016 [\(96\)](#page-39-32), on dietary protein sources and breast cancer risk, where the literature search was conducted up to 2015. Thus, given that a substative period of time $(5 y)$ has elapsed and a number of large-scale prospective studies have been published, especially on red meat, processed meat, and fish, an up-to-date synthesis was urgently required. Moreover, former meta-analyses have rarely assessed the quality of evidence that is important for generating guidelines and recommendations, while we comprehensively assessed the quality of meta-evidence using the NutriGrade scoring system.

As a limitation to the present study it is worth noting that, for some of the food groups, residual components may influence the association with breast cancer: for example, in low-fat versus high-fat dairy, lean meat versus high-fat meat, or low-sugar versus sweetened fruit juices. Since the primary studies did not report results according to fat or sugar contents of food, we could not perform subgroup analyses or conduct the analysis separately according to these factors; therefore, a limitation of this study is that the results may be confounded by a component of the food. For legumes, nuts, and soy, the results were derived from a small number of studies. Most of the studies on these food groups assessed their components, such as fiber, protein, and isoflavones, and so were excluded based on our inclusion criteria. Moreover, meta-evidence for poultry, fish, cereals, and legumes was low; therefore, results for these food groups should be interpreted with caution. Moreover, it should be noted that numerous analyses were performed, and it is conceivable that some associations could be statistically significant as a result of multiple comparisons. Finally, it should be acknowledged that due to the observational design of the primary studies, causality cannot be inferred from these statistical correlations.

This meta-analysis has several strengths; for instance, we are moderately confident in the veracity of the results for most of the food groups, as the primary studies were mostly assessed as being of good quality. The adequate number of included studies also allowed us to conduct multiple

subgroup analyses for important factors, such as menopausal status and presence of estrogen receptor. Additionally, we conducted both linear and nonlinear dose-response analyses, which provide detailed insight into the associations.

In conclusion, the findings of the present meta-analysis show that high intakes of fruits, vegetables, soybeans, and cheese and low intakes of red meat and processed meat are associated with reduced risks of breast cancer. A null association was noted between poultry, fish, egg, fruit juice, dairy, milk (<450 g/day), yogurt, grain, nut, and legume consumption and breast cancer risk, whilst consumption of milk in amounts more than 450 g/day was associated with an increased risk. Finally, it should be acknowledged that causality cannot be inferred from these statistical correlations, indicating the need for further well-conducted RCTs.

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